

the two stations indicates a degree of relationship so slight as to be negligible. During this period the range in the monthly mean values was from 1.933 to 1.907, or 0.026 gr. cal.

On Figures 2 and 3 a cross has been located within the circles that inclose most of the dots, to show the mean value of the solar constant determinations indicated by the dots. A second cross has been located at a point to indicate the mean of solar constant values represented by the dots on the other corresponding figure. These two crosses fall on a line making an angle of very nearly  $45^\circ$  with the axes of ordinates and abscissas, which indicates that the results as summarized above are not inconsistent with Abbot's finding of a high correlation coefficient between the means of groups of synchronous values at Montezuma and Harqua Hala, arranged in order of magnitude of the solar constant determinations at Montezuma.<sup>8</sup> Such a grouping eliminates very largely the day-to-day variations discussed above; and the correlations found are the result of secular changes occurring during the period covered by the three groups, and not real correlations between actual day-to-day changes free from secular variations.

Furthermore, during the period April, 1922, to November, 1924, inclusive, a solar constant value is given on 827 days, and on 299 of these days, or on 1 day out of 2.78, values obtained at both stations were included in the mean for the day. The mean value for the whole period is 1.922. On 36 days the value was less than 1.900, and on 6 of these days, or on only 1 day out of 6, was the value derived from determinations made at

both stations. On 35 days the value was 1.940 or above, and on 4 of these, or on 1 day out of 8.75, the value was derived from determinations made at both stations. Thus, while more than one third of the daily values have been derived from measurements made at both stations, this is the case with only about one seventh of the values that depart from the mean by more than  $\pm 1\%$ .

In view of the above, and for the further reason that one out of every six of these extreme values is by Abbot graded  $U+$  or  $U$  (rather unsatisfactory or unsatisfactory), they are not entitled to as much weight as the more nearly average values.

Abbot holds the view that since there are more *values with large departures* than theory calls for, this is a proof of solar variability.<sup>9</sup> It is not unusual to find such an excess however.<sup>10</sup> Further, in this particular case, it has been shown above that these extreme values have not the same degree of accuracy as the remaining values. Therefore the excess in their number, which is small numerically, can not be accepted as evidence of solar variability.

It seems evident, therefore, that the day-to-day variability of the solar constant determinations, the standard deviation of which is less than  $\pm 0.70$  per cent, depends largely upon whether the solar constant value is derived from determinations made at only one or at both stations; that it reflects unavoidable inaccuracies in pyrheliometric readings, and in extrapolating the readings to zero air mass; and that short-period solar variability, if it exists, falls within the limits of the probable error of the determinations.

<sup>8</sup> Abbot, C. G., *Solar variation and forecasting*, p. 20, Smithsonian Misc. Coll., vol. 77, No. 5.

<sup>9</sup> Solar variability and forecasting. Smithsonian Misc. Coll., vol. 77, No. 5, pp. 16-18.

<sup>10</sup> Brenet, David. *The combination of observations*. (Cambridge, 1917) p. 33.

## THE PROBABLE 24-HOUR TEMPERATURE CHANGE (7 A. M. TO 7 A. M.) AT MONTGOMERY, ALA.

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[Montgomery, Ala., Weather Bureau Office, May 18, 1925]

In this study the probable temperature change for the 24-hour period 7 a. m. to 7 a. m., 90th meridian time, at Montgomery, Ala., has been determined by means of the Gram-Charlier frequency curves for each month of the year, based on 1,000 observations for each month.

The temperature changes were determined from the a. m. observations as recorded on Form No. 1001-Metl., beginning with 1924 and going back far enough to include 1,000 days in each monthly distribution. Each temperature was taken to the nearest even degree before the change was computed, thus giving the change in  $2^\circ$  units. This was done in order to give actual changes considered in the verification of forecasts.

The Gram-Charlier curves were selected because of the relative ease with which they may be computed and their flexibility, which promised good fits in all cases. Reference to Figure 1 shows that very good fits were obtained. It is the belief of the authors that the Gram-Charlier curves are particularly well adapted to meteorological distributions, because of their capacity to take care of both skewness and excess, which are likely to be encountered in appreciable degree, especially in monthly distributions.

It seems unnecessary to give a detailed description of the method of fitting these curves. Reference is made to Fisher,<sup>1</sup> where a lucid explanation may be found, both of the mathematical development and practical applica-

tion. However, for practical application, tables<sup>1</sup> will be necessary containing the values of the generating function

$$\varphi_0(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2} \text{ and its derivatives up to at least}$$

the fourth order.

Computations and the necessary control checks were made for each month, of which only January is shown, in Table 1, to illustrate the method.

Perhaps slightly better fits would have been obtained had the parameters been computed by the method of least squares instead of the method of semi-invariants, but the arithmetical labor entailed would have been almost prohibitive.

It is well known that in dealing with a limited series of observations the third and fourth moments, or in this case the third and fourth semi-invariants, are liable to considerable error, due to the chance presence of a few large departures. It was found in most of the calculations that this error was sufficient to cause a slowing up in the rate of increase near the tails of the curves, or in some cases enough to cause serious secondary inflections. It was found possible to eliminate this undesirable situation by neglecting, in the computation of the third and fourth semi-invariants, the observations at the tails beyond the value  $z = (x - \lambda_1) : \sigma > 4$ . Inasmuch as never more

<sup>1</sup> Fisher, Arne, *Mathematical Theory of Probabilities*, New York, 1923.

<sup>2</sup> Jørgensen, N. R., *Undersøgelser over Frekvensflader og Korrelation*, Copenhagen, 1916.

than five observations were neglected, it appears that there can be no serious objection to such a procedure.

Table 2 shows, by months, the probability of a. m. temperature changes exceeding the stationary limits used in the verification of forecasts. The small chance of a temperature change exceeding 6° in the summer months is noteworthy. In August both the observed frequency

and the calculated frequency of such changes amount to only 8 in 1,000.

Figure 1 shows the observed distribution and the Gram-Charlier curves of best fit, as well as the type equation, and all the constants. An interesting feature is the persistent negative skewness, even in those months when the annual march of temperature is downward. This nega-

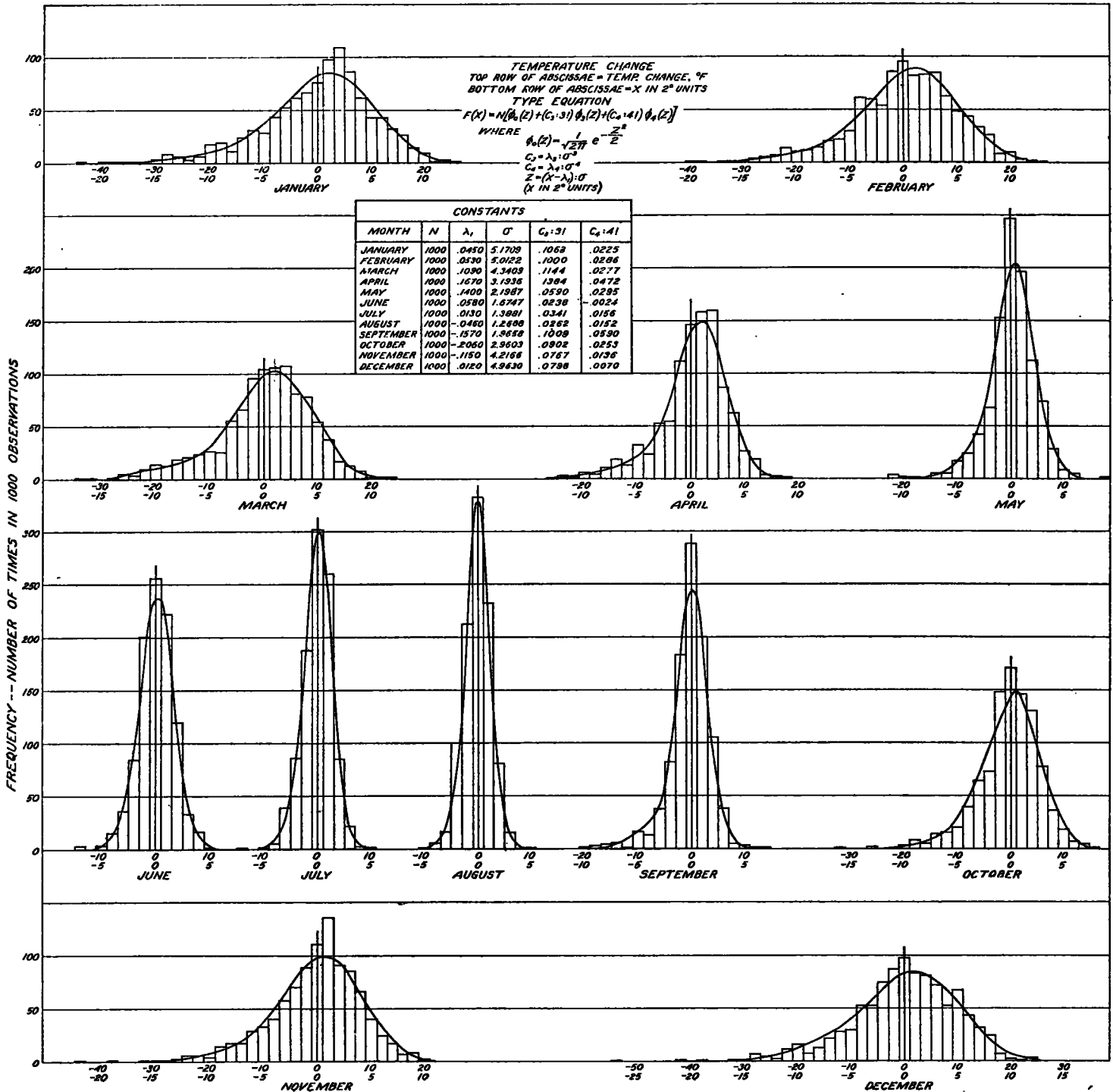


FIG. 1

